

## Price Volatility Modelling – Wheat: GARCH Model Application

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### Abstract

This paper is focused on the modelling of volatility in the agricultural commodity market, specifically on wheat. The aim of this study is to develop an applicable and relevant model of conditional heteroscedasticity from the GARCH family for wheat futures prices. The GARCH (1,1) model has the ability to capture the main characteristics of the commodity market, specifically leptokurtic distribution and volatility clustering. The results show that the forecasted volatility of wheat has a tendency towards standard error reversion in the long-run and the position of price distribution is closed to the normal distribution. The wheat production can be hedged against the price variability with long-term contracts. The price of wheat was influenced during the years of 2005 to 2015 by different events, in particular; financial crisis, increasing grain demand and cross-sectional price variability. The results suggest that agricultural producers should focus on short-term structural events the wheat market, rather than long-term variability.

### Keywords

Price volatility, forecasting, GARCH, wheat price, CME, futures contracts.

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### Introduction

In today's market, commodities are frequently characterized by an increasing trend in the long term. There were significant fluctuations of commodity prices during the period of 2006 to 2009, followed by decreasing prices in the commodity markets due to the financial crisis (Bourdon-Huchet, 2011). In general, high volatility is now evident in every financial market, especially in the case of the commodities market. Examples of this can be found in the problems of storability, seasonality (agricultural products) and price shocks during periods of high volatility. These events are important in the modelling of the chosen commodity.

Volatility modelling is of great interest to many financial analysts and practitioners, in particular, Reider (2009), Zhang (2015), Huang Poon (2005). The origin of the evaluation of price fluctuation is based on price movements and dynamics. Markowitz (1952) firstly focuses on the concept of uncertainty of the asset price. The problem of price variability modelling is primarily due to the difficulties with visibility and the analysis of patterns or other structures of simulated data, see Zhang et al. (2015) The structure of price trajectory has changed over last decade because of the availability

of high-frequency data (Maneessoonthorn, 2015). These movements are often brought on by news announcements or trading activity by institutions. Due to this price variability and volatility, there is an obvious effect on commodity market returns. The ability to accurately and unambiguously forecast and predict volatility in any market has multiple applications. Firstly, there is asset allocation in regards to risk management. There are other utilizations, however the most significant effect of volatility measurement for commodity and equities traders can be felt in the field of portfolio management consisting of assets and derivatives (Hull, 1987). According to a study by Fama (1965), the prices of commodities are characterized with volatile periods changing over time.

There are multiple reasons why to study the price volatility of agricultural commodities. Firstly, the variability of prices is influenced by external shocks or weather. Secondly, we can predict price fluctuation in regards to confidence levels. The volatility is also influenced by the supply of commodities.

The main approach of characteristics of financial time series exhibits the volatility clustering. The papers of Mandelbrot (1963) and Black (1973) documented this evidence in detail

about leptokurtosis and clustering among financial time series. In the last fifteen years, many economists and researchers have begun estimating time series variation by utilizing higher levels of lagged variables. The recent studies of Najand (2002) and Lee, Faff (2009) handle aspects of both volatility clustering and fat-tailed time series. The approach of modern financial volatility modelling was developed by Engle (1982). This paper deals with changing variance using the Autoregressive Conditional Heteroscedasticity (ARCH) model. Later, Bollerslev (1986) introduced the extended version of the ARCH model e.g. with the generalized version – GARCH model. The assumption of this model is that the returns have time-varying conditional variances. According to Val, Pinto and Klotze (2014) the family of GARCH models has significant predictive power when using intra-day data. We assume that the ARCH – family models have a tendency to capture the conditional variance using lags. The conclusion of modelling price volatility is based on the stochastic process. It assumes a financial time series can be defined as a result of a collection of random elements (Douc et al., 2014).

Wheat as a commodity belongs to the family of basic food commodities. The main reason for using wheat is directly connected with nutrition. Wheat has a large impact on mostly agricultural producers, but it influences processors as well.

This paper principally addresses modelling and forecasting the volatility of wheat prices using Generalized Autoregressive Conditional Heteroscedasticity (GARCH). For the purpose of this paper we have selected the time series of wheat futures prices traded on the Chicago Mercantile Exchange (CME). The primary objective of this study is to design an appropriate model of conditional heteroscedasticity from the GARCH family for wheat futures prices based on variance analysis and data. With that in mind, the paper specifically addresses the following research questions:

- i. What are the prediction capabilities of GARCH (1,1) for the CME wheat market?
- ii. Is it possible to use the output of the GARCH model application for the hedging of risks by wheat producers?

### Literature review

Kroner, Kneafsey, Claessenes (1995) studied the long-term forecasting of commodity price volatility. They divided the commodities into different types of forecasts.

In the study by Yang, Haigh, Leatham (2010) there is a GARCH model application under conditions of agricultural liberalization policy. They discovered that the price liberalization caused an increase in the commodity price volatility in the case of several popular commodities – wheat, corn and soybeans. The results included an observation from the 1990s.

Authors Onour, Sergi (2012) employed the competing models with Student t- distribution in the period of 1984 – 2009. The forecast captured the existence of short-term memory behaviour. The paper from Musunuru (2014) focuses on the relationship between wheat and corn in modelling price volatility with the use of the multivariate GARCH model. The results of the paper show that agricultural commodity returns can change significantly over time. Franses, Van Dijk (1996) argue that there are some models, which are not recommended for forecasting such as the Glostten, Jagannathan, and Runkle (GJR) model. The model that is best suited for forecasting non-linear or seasonal time series comes from the GARCH model family. Also, authors such as Tulley and Lucey (2007) have estimated the predictive power of the GARCH model. Baur (2011) employed stochastic volatility models to predict the asymmetry of the volatility of gold. Chkili, Hammoudeh and Nguyen (2014) explored the determinants of change in volatility and forecasting in the example of gas, oil, gold and silver. In a book by Knight and Satchell (2007) there is an evaluation of crucial determinants of commodities, for instance the distribution of contracts, volatility clustering or leverage.

Hansen and Lunde (2005) compare ARCH – type models in the context of describing their conditional variance using exchange rates. According to the authors, the model GARCH (1,1) is unbiased. In addition, Chong (1999) works with modifications of the GARCH model in the stock exchange.

Authors Wei, Wang and Huang (2010) use the non-linear models of the GARCH family to forecast the price of crude oil with the capability to capture the long-term memory over long time periods. Olowe (2009) assumes and advises that the best model for forecasting and evaluating the volatility is GARCH (1,1).

The structure of the paper follows: The next chapter is focused on used methods and data basement. The chapter results and discussion describes results with discussed problems. The last chapter Conclusion summarizes the overall topic according to research questions.

## Materials and methods

The data set consists of 2770 observations. The time series represents the period of 2005 to 2015. The frequency is daily and represents the closing prices of wheat CME Futures. The stationarity is tested using the ADF test (Dickey et al., 1979).

We address the problem of time-variant residual variance by using the ARCH and GARCH models, respectively. The GARCH model is an extended version of the ARCH model that allows for the inclusion of lags of conditional variance.

The ARCH model is extended with the possibility of using the lags of conditional variance. The volatility is dependent on previous observations. Thus the volatility model Generalized Autoregressive Conditional Heteroscedasticity (GARCH) is used (Bollerslev, 1986).

The general form of GARCH(p,q) is:

$$h_t = \alpha_0 + \alpha_1 Y_{t-1}^2 + \alpha_2 Y_{t-2}^2 + \dots + \alpha_p Y_{t-p}^2 + \beta_1 h_{t-1} + \beta_2 h_{t-2} + \dots + \beta_q h_{t-q} + u_t \quad (1)$$

where  $p > 0; q \geq 0; \alpha_0 > 0; \alpha_i \geq 0$  for  $i = 1; 2 \dots p; \beta_y \geq 0$  for  $y = 1; 2 \dots q$  (2)

and  $u_t$  is the an error term.

The value of “p” represents the lags of residual returns and “q” is a lag of variances.

In our application, the lag length choice is based on Akaike Information Criterion. The verification of the model includes the test data for heteroscedasticity and residual autocorrelation. Heteroscedasticity is tested using the ARCH model, i.e. Langrange multiplier for testing to assess the significance of ARCH effects.

ARCH LM test, can be specified as:

$$u_t^2 = \gamma_0 + \sum_{i=1}^q \gamma_i u_{t-i}^2 + v_t \quad (3)$$

With the null hypothesis about the constant conditional variance, i. e.  $H_0: \gamma_i = 0$  for  $i = 1.. q$ ,

and  $v_t$  is the error term.

The serial correlation is tested by using the Breusch-Godfrey test. This test is based on a null hypothesis there is no autocorrelation of order  $p$ .

The Breusch-Godfrey test is based on the following regression.

$$\hat{u}_t = \mu_0 + \mu X_{t,1} + \rho_1 \hat{u}_{t-1} + \dots + \rho_p \hat{u}_{t-p} + \varepsilon_t \quad (4)$$

where  $X$  is a matrix of regressors, and  $\varepsilon_t$  is an error term.

Then the null hypothesis  $H_0$  is  $\rho_1 = \dots \rho_p = 0$

## Results and discussion

Table 1 provides the results of the Augmented Dickey-Fuller test to detect the unit root in time series. The results suggest that the null hypothesis about the unit root in time series cannot be rejected with a 10% significance level. That is, the time series needs to be transformed.

Test statistic with constant	-2.5364
Test statistic with constant and trend	-2.3395

Note: Akaike Information Criterion was used for lag length selection

Source: Own calculation in EViews based on CME data, 2016

Table 1: Augmented Dickey-Fuller test statistic.

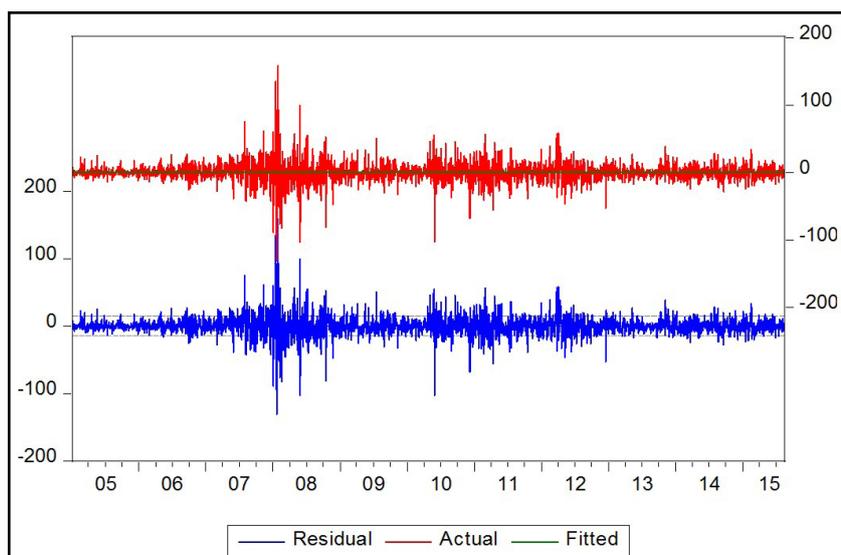
First, the daily closing prices of wheat traded at the CME were transformed using the first differences. Table 2 shows residuals from OLS regression. The table demonstrates that the intercept is not significant with a significance level of 0.05.

Figure 1 displays the residuals from OLS regression and the returns on wheat prices, both actual and fitted ones. The chart captures the volatility clustering and the period with high volatility in 2008 related to the shocks on financial markets. The volatility clustering can be captured by GARCH (p,q). The volatility clustering is based on the assumption

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.059567	0.279947	0.212779	0.8315
R-squared	0.000000	Mean dependent var		0.059567
Adjusted R-squared	0.000000	S.D. dependent var		14.73383
S.E. of regression	14.73383	Akaike info criterion		8.218530
Sum squared resid	601110.2	Schwarz criterion		8.220670
Log likelihood	-11381.66	Hannan-Quinn criter.		8.219303
Durbin-Watson stat	1.889717			

Source: Own calculation in STATA 13 based on CME data, 2016

Table 2: OLS regression.



Source: Own calculation in STATA 13 based on CME data, 2016

Figure 1: Residuals.

that the large values of conditional variance are followed by large values of volatility during the given period.

Figure 1 shows the period with higher volatility during the years from 2005 to 2015. The trend correction occurred at the beginning of 2005, but it did not affect the volatility of residuals. The correction was caused by the growth in grain production, which increased the supply in the market. The period of higher volatility started in 2006. In this case, the grain market paralleled increasing prices of stocks and crude oil (farmdoc, 2006). At the same time, the demand for bio-fuels was increasing as well (Babock and Fabiosa, 2011; Zilberman et al., 2013). The volatility reached its maximum height between 2007 and 2008. The period of 2008 to 2009 can be characterized by financial crisis in the markets and the period of high volatility continued during this time. In this case, the commodities shadowed the stock market with a delay of approximately half of year (CRB, 2013). The recovery of financial and commodity markets occurred in second half of 2009. We can observe the decrease in volatility during this period. At the same time, quantitative easing started in United States and other countries (FRED, 2016; Klotz et al., 2014). Financial markets and the crude-oil market were affected by the so called “Arab spring” in 2013, when the price of crude-oil suddenly increased (Krane, 2015). The other commodities including grains followed this trend, but with a moderate course, so there was not an obvious impact on volatility.

Table 3 and Table 4 contain the parameter estimates of GARCH (1,1) and GARCH (1,0) model, respectively. Both estimated models can be compared according to Akaike information criterion and Schwarz criterion. These criteria indicate how much information has been lost by using the given form of model. Both criteria prefer the use of the GARCH (1,1) model that will be used further.

Then the model is defined as:

$$h_t = 1.106458 + 0.066772 * Y_{t-1}^2 + 0.928826 * h_{t-1} + u_t \quad (5)$$

The serial correlation in GARCH (1,1) model has been tested by using the Breusch-Godfrey Serial Correlation LM test. Table 4 shows that the null hypothesis about no autocorrelation in data can be rejected.

The ARCH LM test has been run to test for heteroscedasticity in the estimated GARCH (1,1) model. The results in Table 5 suggest that the hypothesis about homoscedasticity cannot be rejected, not even at the level of significance  $\alpha = 0.01$ .

Variable	Coefficient	Std. Error	z-Statistic	Prob.
C	-0.081645	0.172681	-0.472808	0.6364
Variance Equation				
C	1.106458	0.153129	7.225650	0.0000
RESID(-1)^2	0.066772	0.005021	13.29797	0.0000
GARCH(-1)	0.928826	0.00506	183.5767	0.0000
R-squared	-0.000092	Mean dependent var		0.059567
Adjusted R-squared	-0.000092	S.D. dependent var		14.73383
S.E. of regression	14.7345	Akaike info criterion		7.704837
Sum squared resid	601165.4	Schwarz criterion		7.713395
Log likelihood	-10667.2	Hannan-Quinn criter.		7.707928
Durbin-Watson stat	1.889544			

Source: Own calculation in STATA 13 based on CME data, 2016

Table 3: - GARCH (1,1).

F-statistic	8.423408	Prob. F(2,2767)		0.0002
Obs*R-squared	16.76302	Prob. Chi-Square(2)		0.0002
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	-6.13E-05	0.279199	-0.00022	0.9998
RESID(-1)	0.058111	0.018983	3.061281	0.0022
R-squared	0.006052	Mean dependent var		-7.55E-15
Adjusted R-squared	0.005333	S.D. dependent var		14.73383
S.E. of regression	14.69449	Akaike info criterion		8.213904
Sum squared resid	597472.5	Schwarz criterion		8.220323
Log likelihood	-11373.26	Hannan-Quinn criter.		8.216222
F-statistic	8.423408	Durbin-Watson stat		2.001479
Prob(F-statistic)	0.000225			

Source: Own calculation in STATA 13 based on CME data, 2016

Table 4: Breusch-Godfrey Serial Correlation LM Test.

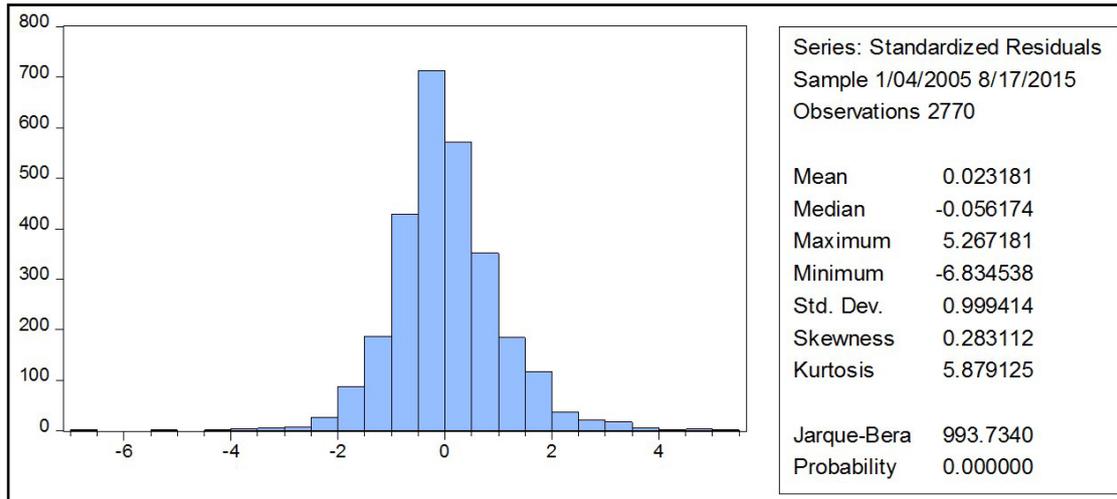
F-statistic	1.727948	Prob. F(2,2767)		0.1888
Obs*R-squared	1.728118	Prob. Chi-Square(2)		0.1887
Variable	Coefficient	Std. Error	t-Statistic	Prob.
C	0.973908	0.046132	21.11134	0
RESID(-1)	0.024982	0.019005	1.314514	0.1888
R-squared	0.000624	Mean dependent var		0.998866
Adjusted R-squared	0.000263	S.D. dependent var		2.212689
S.E. of regression	2.212399	Akaike info criterion		4.426754
Sum squared resid	13543.66	Schwarz criterion		4.431034
Log likelihood	-6126.84	Hannan-Quinn criter.		4.4283
F-statistic	1.727948	Durbin-Watson stat		2.000289
Prob(F-statistic)	0.188782			

Source: Own calculation in STATA 13 based on CME data, 2016

Table 5: - ARCH LM TEST.

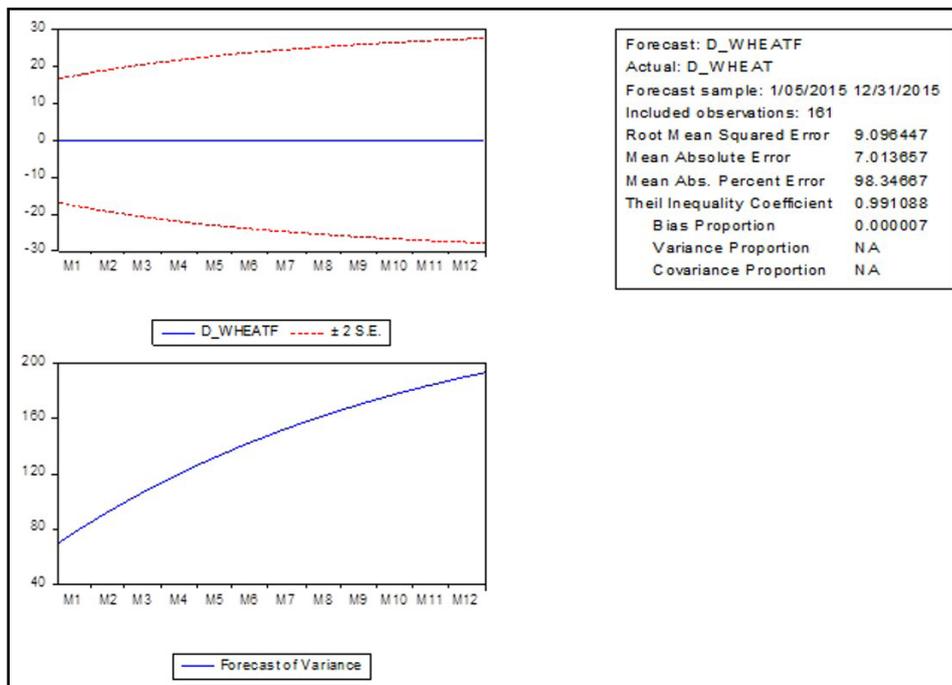
Figure 2 provides the residuals distribution. It suggests that the residuals are not distributed normally. In particular, the visual representation of data shows that the residuals are skewed towards lower values. It can be seen that there is still no

normality. The skewness of data is characterized by the value 0.28. Moreover, the data is characterized by fatter tails with a long peak in the mean. This evidence we can call leptokurtosis.



Source: Own calculation in STATA 13 based on CME data, 2016

Figure 2: Residuals Distribution.



Source: Own calculation in STATA 13 based on CME data, 2016

Figure 3: GARCH (1,1) 2016 static forecast.

Similar results were achieved by the work of Bai et al. (2003). The authors applied the GARCH model, which detected the leptokurtosis within commodity financial time series. On the other hand, there is a study of Zuppiroli, Giha (2015) showing the use of the GARCH model in application to wheat futures prices. The results support the time-varying process of time series similar to volatility behaviour in something like streamflow trajectory. Authors Alberg et al. (2008) recommend that the characteristics of the GARCH model in fat-tailed densities are evidence of forecasting utilization and accuracy. As a result of the research

of the paper by Fang (2008), the GARCH model, which has leptokurtosis, disappeared after introducing the break into the variance equation. There are many studies working with conditional variance that ignore the problem of leptokurtosis and fat-tailed density.

After the verification of the estimated model the prediction for one year has been done. The results are represented in Figure 3. The forecast was made for the year 2016 and included 161 observations. There is a need to take into account that the forecast is based on daily data

for wheat futures contracts traded at the CME.

The results of the static forecast for 2016 suggest that the daily basis of data is more appropriate for short-term predictive purposes. For long-term forecasting, it would be better to use weekly or monthly data. These findings have an implication for agricultural producers in different fields. First, the producers of wheat can hedge their production with short-term futures contracts, due to the ability of a fitted model to predict price fluctuations in short run. Next, the wheat price has a trend to revert to its mean.

The economic implication of wheat volatility modelling has considerable influence on producers and processing decision making. That is, the production cycle is predominantly dependent on external factors. There are also financial participants for instance investors. They are holding large contracts in basic agricultural commodity, as it is case of wheat.

## Conclusion

The aim of this paper was to determine and forecast the volatility of CME daily closing prices of wheat by using stochastic models of conditional heteroscedasticity. The observed data is characterized by the clustering of volatility. This fact is best demonstrated by Figure 1 which shows that the performance of wheat prices recorded the period with high volatility in year 2008. The tests of the fitted model for heteroscedasticity shows that the data's variance was changing over the observed period.

According to Hansen and Lunde (2005) the model GARCH is best utilized when implemented for the purpose of forecasting time series in the financial markets, which are specified by non-constant variance and volatility clustering. Based on that, the authors proposed a GARCH model in order to forecast short-term periods for the daily data of wheat commodity prices traded on the Chicago Mercantile Exchange.

The testing of the fitted model for heteroscedasticity shows that the data has non-constant variance. According to the verification, the GARCH (1,1) model can be used as an appropriate model for wheat futures prices volatility modelling. Based on this model and the static prediction of volatility,

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it is possible to see the convergence of predicted values to the conditional variance in long-term. Thus the GARCH (1,1) is more suitable for short-term predictions. Some other forms of GARCH family models can be considered for long-term volatility of wheat daily closing price prediction. For instance the non-linear models of GARCH family can be tested to predict the volatility with higher statistical significance.

The fitted GARCH model is suitable in the short-term as a tool for risk management when the prediction capability of the model can be used by wheat producers. In the beginning of 2005, there was a trend correction which didn't affect the observed volatility. After that, the increasing grain production on the supply side influenced the level of volatility. During the years 2008 to 2009, wheat prices reached a significant peak in volatility degree. The other reaction in higher wheat price variability was caused by quantitative easing in the United States.

From the economic point of view, the results have wide implications for wheat processing, especially for earnings and agricultural producers. First, the price fluctuation of wheat is more persistent in the short term. That means the fitted model used in this paper is accurate for predictions. The outputs in agriculture are variable all the time, such as natural shocks or weather. Second, there is a problem on the supply side. It means, that the producers cannot respond to the changes of the wheat price in the short term.

The paper also has implications for agricultural producers in hedging techniques. In particular, the wheat producers can sell contracts with longer maturity to protect the price of wheat. This concept can be extended by focusing on the Granger analysis of fundamental events.

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